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| Title | Development of Risk Prediction Equations for Incident Chronic Kidney Disease |
| Type | Article |
| URL | https://clock.uclan.ac.uk/30645/ |
| DOI | https://doi.org/10.1001/jama.2019.17379 |
| Date | 2019 |
| Citation | Nelson, Robert G, Grams, Morgan E, Ballew, Shoshana H, Sang, Yingying, Azizi, Fereidoun, Chadban, Steven J, Chaker, Layal, Dunning, Stephan C, Fox, Caroline et al (2019) Development of Risk Prediction Equations for Incident Chronic Kidney Disease. Journal of the American Medical Association, 322 (21). pp. 2104-2114. ISSN 0098-7484 |
| Creators | Nelson, Robert G, Grams, Morgan E, Ballew, Shoshana H, Sang, Yingying, Azizi, Fereidoun, Chadban, Steven J, Chaker, Layal, Dunning, Stephan C, Fox, Caroline, Hirakawa, Yoshihisa, Iseki, Kunitoshi, Ix, Joachim, Jafar, Tazeen H, Köttgen, Anna, Naimark, David MJ, Ohkubo, Takayoshi, Prescott, Gordon, Rebholz, Casey M, Sabanayagam, Charumathi, Sairenchi, Toshimi, Schöttker, Ben, Shibagaki, Yugo, Tonelli, Marcello, Zhang, Luxia, Gansevoort, Ron T, Matsushita, Kunihiro, Woodward, Mark, Coresh, Josef, Shalev, Varda and for the CKD Prognosis Consortium, ; |

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<https://doi.org/10.1001/jama.2019.17379>

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1 **Development of Risk Prediction Equations for Incident Chronic Kidney Disease**

2

3 **Running title:** Predicting Chronic Kidney Disease

4

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60 **Word counts:** main document=3,535 words

61 2 tables

62 2 figures

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66

67 Key Points:

68 Question: Can development of chronic kidney disease be predicted using readily available demographic,
69 clinical, and laboratory variables?

70 Findings: In this analysis of 5,222,711 individuals in 34 multinational cohorts from 28 countries, 5-year
71 risk prediction equations for CKD were developed and demonstrated high discrimination (median C-
72 statistic for the equation for people without diabetes, 0.85; median C-statistic for the equation for
73 people with diabetes, 0.80) and variable calibration (69% of the study populations had a slope of
74 observed to predicted risk between 0.80 and 1.25). Discrimination and calibration were similar in 9
75 external cohorts consisting of 2,253,540 people.

76 Meaning: Equations for predicting risk of incident chronic kidney disease were developed in over 5
77 million people from 34 multinational cohorts and demonstrated high discrimination and variable
78 calibration in diverse populations.

79 **ABSTRACT**

80 **IMPORTANCE** - Early identification of individuals at elevated risk of developing chronic kidney disease
81 could improve clinical care through enhanced surveillance and better management of underlying health
82 conditions.

83 **OBJECTIVE** – To develop assessment tools to identify individuals at increased risk of chronic kidney
84 disease, defined by reduced estimated glomerular filtration rate (eGFR).

85 **DESIGN, SETTING, AND PARTICIPANTS** – Individual level data analysis of 34 multinational cohorts from
86 the CKD Prognosis Consortium including 5,222,711 individuals from 28 countries. Data were collected
87 from April, 1970 through January, 2017. A two-stage analysis was performed, with each study first
88 analyzed individually and summarized overall using a weighted average. Since clinical variables were
89 often differentially available by diabetes status, models were developed separately within participants
90 with diabetes and without diabetes. Discrimination and calibration were also tested in 9 external
91 cohorts (N=2,253,540).

92 **EXPOSURE** Demographic and clinical factors.

93 **MAIN OUTCOMES AND MEASURES** – Incident eGFR <60 ml/min/1.73 m².

94 **RESULTS** – In 4,441,084 participants without diabetes (mean age, 54 years, 38% female), there were
95 660,856 incident cases of reduced eGFR during a mean follow-up of 4.2 years. In 781,627 participants
96 with diabetes (mean age, 62 years, 13% female), there were 313,646 incident cases during a mean
97 follow-up of 3.9 years. Equations for the 5-year risk of reduced eGFR included age, sex, ethnicity, eGFR,
98 history of cardiovascular disease, ever smoker, hypertension, BMI, and albuminuria. For participants
99 with diabetes, the models also included diabetes medications, hemoglobin A1c, and the interaction
100 between the two. The risk equations had a median C statistic for the 5-year predicted probability of
101 0.845 (25th – 75th percentile, 0.789-0.890) in the cohorts without diabetes and 0.801 (25th – 75th
102 percentile, 0.750-0.819) in the cohorts with diabetes. Calibration analysis showed that 9 out of 13 (69%)

103 study populations had a slope of observed to predicted risk between 0.80 and 1.25. Discrimination was
104 similar in 18 study populations in 9 external validation cohorts; calibration showed that 16 out of 18
105 (89%) had a slope of observed to predicted risk between 0.80 and 1.25.

106 CONCLUSIONS AND RELEVANCE – Equations for predicting risk of incident chronic kidney disease
107 developed in over 5 million people from 34 multinational cohorts demonstrated high discrimination and
108 variable calibration in diverse populations.

109

110 INTRODUCTION

111 Chronic kidney disease (CKD) is a global public health problem that is associated with major adverse
112 health events, including kidney failure, cardiovascular disease, and death. The Global Burden of Disease
113 study estimates that nearly 697 million persons worldwide had reduced estimated glomerular filtration
114 rate (eGFR) or increased albuminuria in 2016, an increase of 70% since 1990.¹ Globally, years of life lost
115 due to CKD increased by 53% in the same period.¹ CKD is the 16th most common cause of years of life
116 lost.² Factors associated with the increased prevalence of CKD include the aging of the population and
117 the increasing prevalence of diabetes, hypertension, and obesity. The ability to identify people at risk for
118 CKD may prevent adverse health outcomes associated with CKD. Moreover, even in those who are
119 diagnosed with CKD, proper management may be hindered by lack of awareness of CKD and its
120 management among clinicians and uncertainties about the underlying risk of CKD progression.

121
122 A kidney failure risk equation may help improve care for patients with established CKD,^{3,4} but relatively
123 little work has been performed to develop predictive tools to identify those at increased risk for
124 *developing* CKD, defined by reduced eGFR, despite the high lifetime risk of CKD, which is estimated to be
125 59.1% in the United States.³ A simple risk assessment tool that helps clinicians quickly identify patients
126 at increased risk of reduced eGFR and provides an estimate of the magnitude of risk for reduced eGFR
127 could lead to better and more targeted surveillance strategies and potentially to better management of
128 the factors associated with reduced eGFR. In the present study, data from multinational cohorts were
129 used to develop and evaluate risk prediction equations for CKD defined by reduced eGFR.

130

131 METHODS

132 This study was approved for use of deidentified data by the institutional review board at the Johns
133 Hopkins Bloomberg School of Public Health, Baltimore, Maryland, USA. The need for informed consent
134 was waived by the institutional review board.

135

136 **Participating cohorts**

137 The Chronic Kidney Disease Prognosis Consortium (CKD-PC) includes study cohorts worldwide that were
138 identified from the general population and from patients at high risk of cardiovascular disease
139 (**eAppendix 1**).⁴⁻⁹ Inclusion criteria required that cohorts included at least 1,000 participants, data on
140 serum creatinine and albuminuria, and 50 or more events of the outcome of interest. Included cohorts
141 consisted of prospective studies, clinical trials, and administrative healthcare datasets. Separate risk
142 models were developed for those with and without diabetes mellitus. The analyses among participants
143 without diabetes included 31 cohorts, and the analyses among participants with diabetes included 15
144 cohorts. Within cohorts, eligible participants were aged ≥ 18 years old with an eGFR >60 ml/min/1.73 m²
145 at baseline. Eligible participants had no previous end-stage kidney disease and had at least one serum
146 creatinine value during follow-up. Because the prevalence and incidence of CKD differ by race/ethnicity,
147 data on race and ethnicity were analyzed from the participating cohorts. Methods used to determine
148 race varied from cohort to cohort, but most cohorts used self-report to define race and ethnicity. Data
149 were collected from April, 1970 through January, 2017.

150

151 **Procedures**

152 The CKD-EPI creatinine equation was used to calculate eGFR.¹⁰ In cohorts where the creatinine
153 measurement was not standardized to isotope dilution mass spectrometry (IDMS), values were
154 multiplied by 0.95 before eGFR calculation.¹¹ We defined diabetes as fasting glucose ≥ 7.0 mmol/L (126
155 mg/dL), non-fasting glucose ≥ 11.1 mmol/L (200 mg/dL), hemoglobin A1c $\geq 6.5\%$, use of glucose lowering

156 drugs, or self-reported diabetes. Hypertension was defined as blood pressure >140/90 mm Hg or the use
157 of anti-hypertensive medications. Smoking was classified as ever smoking vs. never smoking.
158 Participants with a history of myocardial infarction, coronary revascularization, heart failure, or stroke
159 were considered to have a history of cardiovascular disease. Measures of albuminuria were restricted to
160 the urine albumin-to-creatinine ratio. Among participants with diabetes, hemoglobin A1c, oral diabetes
161 medications, and insulin use at baseline were also recorded.

162

163 **Outcomes**

164 The outcome of interest was incident eGFR <60 ml/min/1.73 m². Additional outcomes were eGFR <45
165 ml/min/1.73 m², eGFR <30 ml/min/1.73 m², and 40% decline in eGFR. Participants who developed end-
166 stage kidney disease, mostly identified by procedure codes or by linkage to national registries before a
167 qualifying outpatient level of eGFR were also considered to have developed the outcome of interest. In
168 secondary analyses, we evaluated the risk of confirmed outcomes. Outcomes were defined as confirmed
169 if there were at least three measures of eGFR (one baseline, two during follow-up) and the first eGFR
170 below the threshold was confirmed by a second qualifying eGFR between 90 days and 2 years later, or if
171 the linear slope of eGFR decline crossed the threshold during follow-up (**eAppendix 1**). In both cases, the
172 event date was considered the date of the first qualifying eGFR measurement.

173

174 **Prediction Model Development**

175 The prediction model was built from weighted-average hazard ratios estimated in all participating
176 cohorts and an adjusted baseline risk estimated in cohorts with frequent outcome assessment. To
177 estimate the hazard ratios, each study was first analyzed individually, then combined, weighting the
178 study by the square-root of the number of events in each cohort and capped at 5-times the median
179 study weight. This method was used to ensure that the largest studies did not dominate the analysis due

180 to small within-study variance compared to total variance. We performed complete case analysis,
181 excluding variables which were missing more than 50% of the time in cohort-specific analyses. Since
182 variables were often differentially available by diabetes status (e.g., albuminuria, hemoglobin A1c;
183 missing data shown in **eTable 1A and B**), models were developed separately for participants with
184 diabetes and without diabetes. The primary model included demographic variables (age, sex, ethnicity),
185 eGFR (linear splines with knot at 90 ml/min/1.73 m²), history of cardiovascular disease, ever smoker,
186 hypertension, BMI, and albuminuria. The primary model for participants with diabetes also included
187 diabetes medications (insulin vs. only oral medications vs. none), hemoglobin A1c, and the interaction
188 between the two.

189
190 The albuminuria variable was handled differently for those with vs. without diabetes. For the model
191 among participants with diabetes, missing albuminuria was treated as a dummy variable with reference
192 at a urine albumin-to-creatinine ratio of 10 mg/g. For the model among participants without diabetes,
193 where albuminuria was available only in a minority of individuals, a patch approach was used.¹² Models
194 were fit in all the cohorts using all variables except albuminuria, and data were combined as described
195 above. The weighted average coefficients were then held constant in cohort-specific models among
196 participants with measures of albuminuria to obtain a conditional coefficient for albuminuria, which was
197 then combined for analyses using the weighting described above. This conditional, weighted average
198 coefficient for albuminuria was applied to the observed level of albuminuria less the expected level of
199 albuminuria (**eTable 2**) and combined with the weighted-average coefficients for the other variables in
200 the final model.

201
202 To obtain the adjusted baseline risk for use with the primary model, we held the weighted-average
203 coefficients constant and fit a multivariable competing risk model in the studies with follow-up for

204 mortality and mean time between creatinine measures of less than one year. The adjusted sub-hazard
205 was smoothed using a Weibull distribution and the mean was estimated using weights determined by
206 the method described above. The prediction model then combined the mean adjusted baseline risk with
207 the weighted-average coefficients.

208

209 **Evaluation of Model Performance**

210 To evaluate model discrimination, Harrell's C-statistic was estimated within each cohort and
211 summarized as the median and interquartile range across studies. Model calibration was plotted using
212 observed versus predicted risk per decile of predicted risk at 5 years in each cohort with frequent
213 measures of creatinine (median time between two measurements was approximately 1 year or less and
214 mean follow-up time was at least two years) and quantified using a regression of the deciles of mean
215 observed risk on the mean predicted risk in a zero-intercept linear regression model. Calibration was
216 assessed by visual inspection of the plots (dots showing deciles are close to identity line) and by the
217 slope of observed to predicted risk being near to 1.¹³ To summarize calibration, we determined the
218 number of study populations with an observed risk within 1.25-fold that of the predicted risk (i.e., with a
219 slope between 0.80 and 1.25 (1/0.8)). These metrics of discrimination and calibration were also
220 calculated within 9 external validation cohorts selected from OptumLabs® Data Warehouse. **eAppendix**
221 **1** describes the methods for selecting centers for the nine external validation cohorts. The OptumLabs
222 Data Warehouse contains deidentified longitudinal health information on patients receiving care in
223 health systems participating in the OptumLabs collaborative research and innovation center in the U.S.
224 The database includes people ages 18 to 88 years, from diverse ethnicities and geographical regions
225 across the United States (**eTable 3**). The electronic health record (EHR)-derived data include a subset of
226 EHR data that have been normalized and standardized across health systems into a single database,
227 including information on demographics, laboratory values, encounter and discharge codes.¹⁴

228
229 To compare the newly developed models to existing equations, predicted risks using the newly
230 developed models were compared with risks calculated using two published equations identified in a
231 recent review¹⁵ (herein referred to as the Chien equation¹⁶ and the O'Seaghdha equation¹⁷, respectively
232 **eAppendix 4**). The Chien equation was developed in 5,168 Chinese individuals who underwent baseline
233 health examinations at the National Taiwan University Hospital¹⁶ and annual follow-up examinations
234 that included measurements of serum creatinine concentration for assessing the outcome of reduced
235 eGFR. During a median follow-up of 2.2 years, 190 individuals developed CKD. We used the Chien clinical
236 equation, which included age, body mass index, diastolic blood pressure, and history of type 2 diabetes
237 and stroke. The O'Seaghdha prediction model was developed in the predominantly white population of
238 Framingham, Massachusetts, using baseline serum creatinine and a subsequent measure 10 years later.
239 Among the 2,490 individuals aged 45-64 years included in this study, 229 developed eGFR <60
240 ml/min/1.73m² at 10 years. The O'Seaghdha model included age, hypertension, diabetes, eGFR
241 category, and albuminuria.¹⁷

242
243 The performance of the newly developed model, the Chien equation, and the O'Seaghdha equation
244 were compared in the CKD-PC cohorts that provided individual-level participant data and had the
245 required variables for all equations. Differences in C-statistics were estimated within all cohorts and
246 then summarized using random-effects meta-analysis. Brier scores, the mean squared difference
247 between the predicted risk vs observed binary outcomes, were used to evaluate which risk equation
248 showed the best calibration within each cohort (**eAppendix 4**).¹⁸ Brier scores were assessed only within
249 the subset of cohorts with frequent assessments of creatinine. Comparisons of the discrimination and
250 calibration were also performed within the 9 external validation cohorts from OptumLabs Data
251 Warehouse.

252
253 All analyses were performed in Stata 15 (StataCorp. 2017. College Station, TX: StataCorp LLC). Statistical
254 significance was determined using a two-sided test with a threshold p-value of <0.05.

255

256 **RESULTS**

257 Overall, 5,222,711 participants were included (**eTable 4**), of whom 781,627 (15.0%) had diabetes.

258 Baseline characteristics of participants in the 34 individual cohorts are shown in **Table 1** according to the
259 presence or absence of diabetes. The population without diabetes had a mean age of 54 years (SD, 16)
260 and 38% were female. The population with diabetes had a mean age of 62 years (SD, 11) and 13% were
261 female, owing primarily to the Veterans Administration cohort, which was 97% male.

262

263 Among the 4,441,084 participants without diabetes, there were 660,856 (14.9%) incident cases of eGFR
264 <60 ml/min/1.73m² during a mean follow-up of 4.2 years, and 374,513 (56.7%) of them were confirmed
265 by subsequent eGFR measurements. Among the 781,627 participants with diabetes, there were 313,646
266 (40.1%) incident cases during a mean follow-up of 3.9 years, and 212,246 (67.7%) of them were
267 confirmed by subsequent eGFR measurements. The number of participants and the total and confirmed
268 number of events of incident reduced eGFR in the nondiabetic and diabetic cohorts are shown in **eTable**
269 **5**.

270

271 **Risk factors for reduced eGFR**

272 Weighted-average sub-hazard ratios of major risk factors for incident eGFR <60 ml/min/1.73m² are
273 shown in **Table 2** and for other eGFR thresholds in **eTable 6** according to the presence or absence of
274 diabetes. Older age, female sex, black race, hypertension, history of cardiovascular disease, lower eGFR,
275 and higher urine albumin-to-creatinine ratio were each significantly associated with incident eGFR <60

276 ml/min/1.73m² in both the diabetic and nondiabetic cohorts. Smoking was significantly associated with
277 incident eGFR <60 ml/min/1.73m² only in the nondiabetic cohorts, and elevated hemoglobin A1c and
278 presence and type of diabetes medicines were significantly associated with incident eGFR <60
279 ml/min/1.73m² in the diabetic cohorts.

280

281 **Discrimination**

282 Measures of discrimination for the 5-year predicted probability of incident eGFR <60 ml/min/1.73m²,
283 based on the predictive models, are shown separately for the nondiabetic and diabetic cohorts in **eTable**
284 **7A**. The median C statistic for the 5-year predicted probability of all eGFR events <60 ml/min/1.73m²
285 was 0.845 (25th – 75th percentile, 0.789-0.890) in the cohorts without diabetes and 0.801 (25th – 75th
286 percentile, 0.750-0.819) in the cohorts with diabetes, reflecting good discrimination. For confirmed
287 eGFR events <60 ml/min/1.73m², the median C statistic was 0.869 (25th – 75th percentile, 0.823-0.897) in
288 the cohorts without diabetes and 0.808 (25th – 75th percentile, 0.794-0.836) in the cohorts with diabetes.
289 Measures of discrimination for the lower incident eGFR thresholds are shown in **eTable 7B**.

290

291 **Predicted absolute risk**

292 Adjusted baseline sub-hazards for eGFR <60 ml/min/1.73m² were computed over time in nondiabetic
293 and diabetic cohorts with frequent measures of creatinine using baseline covariates from the cohorts
294 and weighted-average coefficients from the models (**Figure 1**). The figure illustrates the variability in the
295 adjusted absolute risk across the cohorts that was unexplained by the covariates included in the models.
296 Similar findings are shown for the lower incident eGFR thresholds in **eFigure 1** for the nondiabetic
297 cohorts and **eFigure 2** for the diabetic cohorts.

298

299 Equations for the 5-year predicted risk of incident eGFR<60 ml/min/1.73m², based on the predictive
300 models and the mean baseline sub-hazards, are shown separately for individuals with or without
301 diabetes in **eTable 8** and are available online at <http://ckdpcrisk.org/ckdrisk>. The predicted 5-year
302 absolute risk of incident eGFR<60 ml/min/1.73m² in individuals without and with diabetes at three ages
303 and for various combinations of risk factors are shown in **Figure 2** and in greater detail for all three
304 incident eGFR thresholds in **eTables 9** and **10**. A wide range of risk was seen, and the level of risk was
305 strongly associated with the demographic features and co-morbid conditions. The absolute risk was
306 generally higher in persons with diabetes than in those without and increased with age regardless of the
307 presence or absence of diabetes. Elevated albuminuria was also significantly associated with the
308 absolute risk regardless of the presence or absence of diabetes. The 5-year absolute risk for confirmed
309 eGFR reduction followed the same pattern as for the unconfirmed endpoint, with lower absolute risk for
310 the confirmed endpoints (**eTables 9** and **10**). Equations for the 5-year predicted risk of other outcomes
311 are shown in **eTables 11** and **12**.

312

313 **Calibration**

314 Model calibration was assessed visually by plotting observed versus predicted risk per decile of
315 predicted risk at 5 years in the cohorts with frequent measures of creatinine. Plots for the eGFR <60
316 ml/min/1.73m² endpoint are shown in **eFigure 3** and for the lower eGFR endpoints in **eFigures 4** and **5**.
317 The plots reflected the performance of the equations for the primary endpoint in the cohorts, with 9 of
318 the 13 (69%) study populations showing a slope of observed to predicted risk between 0.80 and 1.25
319 (**eTable 13**). Calibration was generally better for the eGFR <60 ml/min/1.73m² endpoint compared to
320 the lower eGFR endpoints, where it was poor in some cohorts (**eTables 14-15**). For example, for eGFR
321 <45 ml/min/1.73 m², just 5 of 13 (38%) study populations showed a slope between 0.80 and 1.25. For

322 eGFR <30 ml/min/1.73 m², just 4 out of 11 (36%) study populations showed a slope between 0.80 and
323 1.25. Calibration, by design, was best in the development cohorts with the highest number of events.

324

325 **External validation**

326 Model discrimination was tested in 18 study populations in 9 external validation cohorts (N=2,253,540,
327 **eTable 16**). There were 288,462 events over 4.1 years of follow-up in the population without diabetes
328 and 78,697 events over 3.5 years of follow-up in the population with diabetes. Discrimination was
329 similar to that observed in the development cohorts. The median C statistic for the 5-year predicted
330 probability of all eGFR events <60 ml/min/1.73m² was 0.84 (25th – 75th percentile, 0.83-0.87) in the
331 population without diabetes and 0.81 (25th – 75th percentile, 0.80-0.82) in the population with diabetes
332 (**eTable 17**). Calibration analysis showed that 16 out of 18 (89%) study populations with a slope between
333 0.80 and 1.25 (**eFigure 6, eTable 18**). Discrimination and calibration for the lower eGFR endpoints are
334 shown in **eFigures 7-8** and **eTables 17-18**. For example, for eGFR <45 ml/min/1.73 m², 15 out of 18
335 (83%) of study populations showed a slope between 0.80 and 1.25. For eGFR <30 ml/min/1.73 m², 11
336 out of 18 (61%) study populations showed a slope between 0.80 and 1.25. Differences in calibration
337 could not be explained by differences in mean baseline characteristics in the underlying study
338 populations.

339

340 **Comparison to existing equations**

341 The newly developed model for eGFR <60 ml/min/1.73m² in the absence of diabetes had better
342 discrimination than the Chien equation (random-effects meta-analyzed difference in C statistic, 0.094,
343 95% CI: 0.071-0.117) and the O'Seaghda equation (random-effects meta-analyzed difference in C
344 statistics, 0.020, 95% CI: 0.015-0.025) when compared in the CKD-PC cohorts. Similarly, the Brier score
345 was lower using the newly developed equation in the cohorts with frequent measures of creatinine,

346 indicating superior calibration for the newly developed equation (**eTable 19**). In the presence of
347 diabetes, the newly developed model had better discrimination than the Chien equation (random-
348 effects meta-analyzed difference in C statistic, 0.107, 95% CI: 0.087-0.128) and the O'Seaghdha equation
349 (random-effects meta-analyzed difference in C statistics, 0.037, 95% CI: 0.030-0.044) and lower Brier
350 scores in two out of three cohorts with frequent measures of creatinine. When evaluated in the 9
351 external validation cohorts, model discrimination and calibration were also better using the newly
352 developed equations compared to the Chien and O'Seaghdha equations (**eTable 20**).

353

354 **DISCUSSION**

355 Risk prediction models were developed that facilitated prediction of the 5-year probability of reduced
356 eGFR in diverse populations of men and women with variable ages and ethnicity. Models were
357 developed separately for people with vs. without diabetes. Readily available demographic, clinical, and
358 laboratory variables were used in these risk models, so that risk calculators from these models could
359 conceivably be added to electronic health records to identify patients at increased risk for developing
360 reduced eGFR. Further study is needed to determine whether these risk equations can improve care. For
361 example, future study could assess whether focusing resources on patients at highest risk of developing
362 chronic kidney disease improves blood pressure control and/or weight loss. Future study might also
363 determine whether prescribing medications to improve albuminuria or control diabetes might prevent
364 occurrence of reduced eGFR in those at risk.

365

366 Several prediction models of CKD exist for use in the general population.^{16,17,19,20} Equations previously
367 developed to identify people at risk for incident eGFR <60 ml/min/1.73m² included the Chien equation
368 and the O'Seaghdha equation, both of which have been externally validated.¹⁵⁻¹⁷ External validation of
369 the Chien clinical model was previously done in 3,205 Chinese adults from the Chin-Shan Community

370 Cardiovascular Cohort. Moderate discrimination was observed for the clinical prediction model in the
371 development cohort (c-statistic = 0.77), but the discriminatory power of the model was greatly reduced
372 in the external validation cohort (c-statistic = 0.67).¹⁶ The O'Seaghdha risk score was validated in 1,777
373 individuals from the ARIC study (c-statistic = 0.79 in Framingham and 0.74 in ARIC).¹⁷ These prior studies
374 did not develop separate equations for those with vs. without diabetes. The present study, which
375 developed scores separately for people with vs. without diabetes, demonstrated higher C-statistics and
376 better calibration than both the clinical Chien and the O'Seaghdha equations. This was true in the CKD-
377 PC cohorts used in development of the equations as well as in the 9 external validation cohorts.

378

379 Risk prediction models that estimate the absolute risk of specific adverse health outcomes have become
380 increasingly popular clinical decision-making tools in recent years, and novel approaches to analyzing
381 existing data are emerging that may enhance prediction.²¹ Several models have been developed for
382 estimating the risk of prevalent and incident CKD and end-stage kidney disease,^{4,16,17,19,20,22-24} but even
383 those with good discriminative performance have not always performed well in cohorts of people
384 outside the original derivation cohort.¹⁵ In our study, we show that the incidence of low eGFR varies
385 across settings, even after adjustment for variable distribution of risk factors, providing an explanation
386 for differences in calibration in prior studies.

387

388 Calibration is an essential aspect of risk prediction, particularly when absolute risk thresholds are used
389 to drive clinical care. A tool that overestimates risk may result in unnecessary treatment, whereas one
390 that underestimates risk may delay optimal management. By design, calibration in the development
391 cohorts in our study was set to the overall weighted risk. Hence, we focused on calibration on external
392 cohorts for an unbiased assessment. Surprisingly, in external validation in over 2 million people, model
393 calibration was even better than that in the development cohorts, suggesting that it may generalize well

394 to US electronic health systems like those represented in OptumLabs Data Warehouse. Other strengths
395 of this study include the large sample sizes of the nondiabetic and diabetic cohorts, and the broad
396 clinical, geographic, and ethnic diversity of the individuals in those cohorts. However, we note that
397 calibration of the developed risk equations may be poor in populations that differ substantially in the
398 adjusted incidence of reduced eGFR or in which ascertainment of reduced eGFR is more or less sensitive.
399

400 **Limitations**

401 This study has several limitations. First, the absence of albuminuria data in most nondiabetic cohorts
402 included in this study required that a statistical patch derived from nondiabetic cohorts with
403 albuminuria data be applied to the remaining cohorts in order to estimate how inclusion of albuminuria
404 altered the models. This approach allows valid estimation of risk even in the absence of albuminuria,
405 although clinical assessment of albuminuria improved risk estimation and detection of early stage CKD
406 defined by elevated albuminuria (A-stages) in the absence of reduced kidney function (G stages 1-2).²⁵
407 Second, the risk equations developed in this study incorporated routinely collected demographic,
408 clinical, and laboratory data and their predictive accuracy might be enhanced by incorporating other
409 variables, including genotype data or newly identified biomarkers of early CKD.²⁶ Third, the risk
410 prediction equations developed in this study were intended to identify persons at increased risk of an
411 intermediate health outcome. The risks of progression from CKD to kidney failure, cardiovascular
412 disease, or death were not assessed by these equations. Fourth, no minimum change in eGFR was
413 required in the primary predictive model to become a case of CKD, so someone with a baseline eGFR of
414 61 ml/min/1.73m² and a follow-up eGFR of 59 ml/min/1.73m² would be considered to have the
415 outcome of interest. Fifth, calibration varied across setting, with particularly poor performance in some
416 of the research cohorts. The models for eGFR <45 and eGFR<30 ml/min/1.73 m² were poorly calibrated

417 in many of the development cohorts, which may be due in part to the low number of events and
418 relatively short follow-up time.

419

420 **Conclusions**

421 Equations for predicting risk of incident chronic kidney disease were developed in over 5 million people
422 from 34 multinational cohorts and demonstrated high discrimination and variable calibration in diverse
423 populations.

424

425 **Contributors:** MEG and JC had full access to all the data in the study and take responsibility for the
426 integrity of the data and the accuracy of the data analysis. RGN, JC, RG, MEG, KM, MW, and VS were
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431 integrity of the work. The corresponding author attests that all listed authors meet authorship criteria
432 and that no others meeting the criteria have been omitted.

433

434 **Funding:** AK was supported by grant KO 3598/5-1 of the German Research Foundation. The CKD
435 Prognosis Consortium (CKD-PC) Data Coordinating Center is funded in part by a program grant from the
436 US National Kidney Foundation (NKF funding sources include Boehringer Ingelheim) and the National
437 Institute of Diabetes and Digestive and Kidney Diseases (R01DK100446). A variety of sources have
438 supported enrollment and data collection including laboratory measurements, and follow-up in the
439 collaborating cohorts of the CKD-PC (**eAppendix 3**). These funding sources include government agencies
440 such as national institutes of health and medical research councils as well as foundations and industry
441 sponsors. The funders of the study had no role in the design and conduct of the study; collection,
442 management, analysis, and interpretation of the data; preparation, review, or approval of the
443 manuscript; and decision to submit the manuscript for publication. In addition, the funders had no right
444 to veto publication or to control the decision regarding to which journal the paper would be submitted.

445

446 **Competing interests:** All authors have completed the ICMJE uniform disclosure form at
447 www.icmje.org/coi_disclosure.pdf (available on request from the corresponding author) and declare: no
448 support from any organization for the submitted work; no financial relationships with any organizations

449 that might have an interest in the submitted work in the previous three years; no other relationships or
450 activities that could appear to have influenced the submitted work.

451

452 **Data sharing:** CKD-PC has agreed with collaborating cohorts not to share data outside the consortium.

453 Each participating cohort has its own policy for data sharing.

454

455 **ACKNOWLEDGMENTS**

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457 Supplement; most cohorts received a small [$< \$2000$] amount of funds to offset the costs of data

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Table 1. Baseline characteristics of the participants in the 31 nondiabetic and 15 diabetic cohorts.

| Study | Country | N | Age | Female | eGFR (ml/min/1.73m ²) | History of CVD | Hypertension | Smoking | BMI |
|---------------------|-------------|--------|---------|--------------|--------------------------------------|-------------------|--------------|-----------------|--------|
| Nondiabetic cohorts | | | | | | | | | |
| ARIC | USA | 12757 | 54 (6) | 7082 (56%) | 103 (14) | 980 (8%) | 4437 (35%) | 7367 (58%) | 27 (5) |
| AusDiab | Australia | 6281 | 50 (12) | 3471 (55%) | 88 (14) | 306 (5%) | 1580 (25%) | 2528 (41%) | 27 (5) |
| Beijing | China | 948 | 59 (9) | 496 (52%) | 85 (12) | 127 (13%) | 363 (38%) | 321 (34%) | 25 (3) |
| CARE | Canada | 2923 | 57 (9) | 343 (12%) | 80 (13) | 2923 (100%) | 2432 (83%) | 2332 (80%) | 28 (7) |
| CHS | USA | 2170 | 73 (4) | 1341 (62%) | 77 (11) | 409 (19%) | 1280 (59%) | 1122 (53%) | 27 (5) |
| CIRCS | Japan | 10022 | 54 (9) | 6275 (63%) | 90 (14) | 97 (1%) | 3353 (33%) | 3507 (35%) | 23 (3) |
| ESTHER | Germany | 3394 | 61 (6) | 1885 (56%) | 92 (15) | 458 (13%) | 2213 (65%) | 1548 (47%) | 27 (4) |
| Framingham | USA | 2353 | 58 (9) | 1290 (55%) | 91 (16) | 180 (8%) | 828 (35%) | 368 (16%) | 28 (5) |
| Geisinger | USA | 229448 | 50 (16) | 132677 (58%) | 95 (18) | 23403 (10%) | 113953 (50%) | 110640 (49%) | 30 (7) |
| GLOMMS 2 | UK | 24321 | 61 (14) | 13598 (56%) | 81 (15) | 1962 (8%) | 910 (4%) | NA | NA |
| Gubbio | Italy | 1249 | 54 (6) | 714 (57%) | 85 (11) | 44 (4%) | 443 (35%) | 688 (55%) | 28 (4) |
| HUNT | Norway | 34430 | 46 (13) | 19114 (56%) | 102 (15) | 1170 (3%) | 12377 (36%) | 17992 (53%) | 26 (4) |
| IPHS | Japan | 70557 | 60 (10) | 47934 (68%) | 86 (12) | 3603 (5%) | 33626 (48%) | 19565 (28%) | 23 (3) |
| JHS | USA | 2164 | 48 (11) | 1312 (61%) | 102 (17) | 94 (4%) | 885 (41%) | 596 (28%) | 31 (7) |
| JSHC | China | 461797 | 63 (8) | 279934 (61%) | 94 (11) | 34567 (9%) | 193996 (42%) | 62947 (14%) | 23 (3) |
| Maccabi | Israel | 939309 | 43 (15) | 546440 (58%) | 104 (17) | 55138 (6%) | 213398 (23%) | 231695 (25%) | 27 (5) |
| MESA | USA | 4954 | 61 (10) | 2623 (53%) | 86 (13) | 1 (0%) | 2051 (41%) | 2600 (53%) | 28 (5) |
| Mt Sinai BioMe | USA | 14590 | 48 (14) | 8998 (62%) | 93 (19) | 722 (5%) | 6385 (44%) | 3910 (28%) | 29 (7) |
| Ohasama | Japan | 2346 | 60 (10) | 1483 (63%) | 98 (11) | 91 (4%) | 832 (35%) | 349 (19%) | 24 (3) |
| Okinawa8393 | Japan | 1624 | 50 (10) | 957 (59%) | 100 (13) | 0 (0%) | NA | NA | 24 (3) |
| Pima | USA | 2733 | 28 (11) | 1626 (59%) | 125 (13) | NA | 272 (10%) | 793 (47%) | 33 (8) |
| PREVEND | Netherlands | 5977 | 49 (12) | 3057 (51%) | 97 (14) | 247 (4%) | 1773 (30%) | 4160 (70%) | 26 (4) |

| | | | | | | | | | |
|------------------|-------------|---------|---------|---------------|----------|--------------|---------------|--------------|--------|
| Rancho Bernardo | USA | 639 | 64 (10) | 369 (58%) | 75 (11) | 49 (8%) | 232 (36%) | 354 (56%) | 26 (4) |
| RCAV | USA | 1765629 | 59 (13) | 133822 (8%) | 85 (15) | 256353 (15%) | 1196576 (68%) | NA | 29 (6) |
| RSIII | Netherlands | 2292 | 56 (6) | 1333 (58%) | 87 (12) | 126 (5%) | 1375 (60%) | 1572 (69%) | 27 (4) |
| SCREAM | Sweden | 716952 | 52 (17) | 392827 (55%) | 95 (17) | 40554 (6%) | 177249 (25%) | NA | NA |
| SEED | Singapore | 2358 | 54 (9) | 1246 (53%) | 88 (14) | 156 (7%) | 1164 (50%) | 700 (30%) | 26 (4) |
| Taiwan MJ | Taiwan | 101216 | 41 (12) | 52658 (52%) | 91 (15) | 2474 (2%) | 16560 (16%) | 26037 (28%) | 23 (3) |
| TLGS | Iran | 8502 | 37 (13) | 4753 (56%) | 81 (13) | 171 (2%) | 1404 (17%) | 1839 (22%) | 26 (5) |
| Tromso | Norway | 6007 | 58 (10) | 3522 (59%) | 95 (12) | 283 (5%) | 3183 (53%) | 3877 (65%) | 26 (4) |
| ULSAM | Sweden | 1142 | 50 (1) | 0 (0%) | 98 (10) | 5 (0%) | 416 (36%) | NA | 25 (3) |
| | | 4441084 | 54 (16) | 1673180 (38%) | 93 (17) | 426693 (10%) | 1996070 (45%) | 509588 (26%) | 27 (6) |
| Diabetic cohorts | | | | | | | | | |
| ADVANCE | Multiple* | 9339 | 66 (6) | 3774 (40%) | 83 (13) | 2235 (24%) | 8003 (86%) | 4024 (43%) | 28 (5) |
| AusDiab | Australia | 427 | 59 (11) | 189 (44%) | 84 (13) | 70 (16%) | 287 (67%) | 205 (48%) | 30 (6) |
| Beijing | China | 343 | 62 (9) | 168 (49%) | 85 (12) | 80 (23%) | 184 (54%) | 127 (37%) | 25 (4) |
| Geisinger | USA | 34463 | 58 (15) | 16842 (49%) | 93 (18) | 8606 (25%) | 27251 (79%) | 17563 (52%) | 34 (8) |
| HUNT | Norway | 1564 | 54 (12) | 709 (45%) | 95 (14) | 130 (8%) | 932 (60%) | 892 (57%) | 28 (5) |
| JHS | USA | 390 | 54 (10) | 241 (62%) | 101 (18) | 46 (12%) | 310 (79%) | 131 (34%) | 35 (8) |
| Maccabi | Israel | 72480 | 60 (13) | 32972 (45%) | 92 (15) | 18147 (25%) | 54586 (75%) | 21733 (30%) | 31 (6) |
| MESA | USA | 659 | 63 (9) | 304 (46%) | 90 (15) | 0 (0%) | 455 (69%) | 343 (52%) | 31 (6) |
| Mt Sinai BioMe | USA | 2652 | 54 (13) | 1598 (60%) | 91 (19) | 511 (19%) | 2013 (76%) | 923 (37%) | 32 (8) |
| NZDCS | New Zealand | 14819 | 58 (13) | 7152 (48%) | 86 (16) | 2260 (15%) | 10197 (82%) | 6469 (44%) | 32 (7) |
| Pima | USA | 933 | 43 (14) | 577 (62%) | 114 (17) | NA | 335 (36%) | 291 (40%) | 34 (8) |
| RCAV | USA | 607132 | 63 (10) | 20241 (3%) | 83 (15) | 157611 (26%) | 551356 (91%) | NA | 32 (6) |
| SCREAM | Sweden | 34307 | 60 (16) | 14224 (41%) | 91 (17) | 8041 (23%) | 20408 (59%) | NA | NA |
| SEED | Singapore | 1029 | 58 (9) | 508 (49%) | 88 (15) | 151 (15%) | 742 (72%) | 311 (30%) | 28 (5) |

| | | | | | | | | | |
|--------|-------------|--------|---------|--------------|---------|--------------|--------------|-------------|--------|
| ZODIAC | Netherlands | 1090 | 63 (11) | 522 (48%) | 77 (12) | 310 (28%) | 794 (73%) | 249 (23%) | 29 (5) |
| | | 781627 | 62 (11) | 100021 (13%) | 85 (15) | 198198 (25%) | 677853 (87%) | 53261 (38%) | 32 (6) |

Values are mean (SD) or percent of total N. Abbreviations: BMI, body mass index; CVD, cardiovascular disease; eGFR, estimated glomerular filtration rate; NA, not available. Racial distributions of the cohorts are available in **eTable 4** and the citations for each study are available in **eAppendix 2**.

* Participants are from Australia, Canada, China, Czech Republic, Estonia, France, Germany, Hungary, India, Ireland, Italy, Lithuania, Malaysia, Netherlands, New Zealand, Philippines, Poland, Russia, Slovakia, and United Kingdom.

Table 2. Weighted-average sub-hazard ratios of major risk factors for incident eGFR<60 ml/min/1.73m² in the nondiabetic and diabetic cohorts.

| Risk factors | Sub-Hazard Ratios (95% CI) for Incident eGFR<60ml/min/1.73m ² | |
|---|--|-------------------|
| | Non-diabetic model | Diabetic model |
| Age, per 5y | 1.29 (1.27, 1.32) | 1.14 (1.13, 1.15) |
| Female | 1.20 (1.18, 1.22) | 1.15 (1.11, 1.18) |
| Black | 1.20 (1.13, 1.27) | 1.10 (1.02, 1.18) |
| eGFR 60-90, per -5 ml | 1.58 (1.57, 1.59) | 1.43 (1.41, 1.44) |
| eGFR 90+, per -5 ml | 1.37 (1.34, 1.41) | 1.16 (1.14, 1.19) |
| History of CVD | 1.22 (1.18, 1.26) | 1.21 (1.17, 1.24) |
| Ever smoker | 1.13 (1.10, 1.16) | 1.00 (0.96, 1.04) |
| Hypertension | 1.43 (1.40, 1.46) | 1.44 (1.39, 1.50) |
| BMI, per 5 kg/m ² | 1.07 (1.05, 1.08) | 1.05 (1.04, 1.07) |
| ACR, per 10-fold increase | 1.42 (1.37, 1.48) [†] | 1.45 (1.42, 1.49) |
| HbA1c (for oral DM meds), per 1% | | 1.06 (1.05, 1.07) |
| Insulin vs. oral DM meds (at 7% hba1c) | | 1.11 (1.05, 1.19) |
| No meds vs. oral DM meds (at 7% hba1c) | | 0.86 (0.83, 0.89) |
| Interaction: HbA1c * insulin vs. oral DM meds, per 1% | | 1.02 (1.00, 1.05) |
| Interaction: HbA1c * No meds vs. oral DM meds, per 1% | | 1.04 (1.02, 1.06) |
| ACR missing indicator (set ACR=10) | | 0.96 (0.93, 1.00) |

[†]ACR was modeled using a patch in the non-diabetes model in which the coefficient for ACR was estimated in the population with available ACR with the other coefficients fixed. The model allows for prediction when ACR is missing.

eTables 9 and 10 provide absolute risk and risk difference scenarios.

Abbreviations: ACR, urine albumin-creatinine ratio; BMI, body mass index; CVD, cardiovascular disease; eGFR, estimated glomerular filtration rate; HbA1c, hemoglobin A1c.

Figure 1. Variation in baseline adjusted competing risk of incident eGFR<60 ml/min/1.73m² in nondiabetic (A and C) and diabetic (B and D) cohorts with frequent measures of serum creatinine concentration. All events (confirmed and unconfirmed) are shown in Panels A and B and confirmed events are shown in Panels C and D.

Numbers after the cohort name in the key indicate the mean follow-up time in years. Each line represents the adjusted baseline risk in an individual cohort. The risk was determined by holding the weighted-average coefficients constant and fitting a multivariable competing risk model in each study. The adjusted sub-hazard was smoothed using a Weibull distribution. The pooled line represents the weighted mean which is used in the prediction equation.

Figure 2. Predicted 5-year absolute risk of incident eGFR <60 ml/min/1.73m² is shown for various scenarios in three ages and albuminuria categories in nondiabetic and diabetic individuals. All 5-year risks were computed for hypothetical individuals with a baseline eGFR of 90 ml/min/1.73m². For the 5-year predicted risk in a hypothetical individual with diabetes, the hemoglobin A1c was also set to 7.7% and the individual was assumed to be receiving an oral diabetes medicine. Scenarios: Sex: male/female, Ethnicity: non-black/black, History of CVD: yes/no, Smoker: yes/no, Hypertension: yes/no, BMI: 25/35 kg/m², ACR: not available (N/A; equation without ACR)/50/500 mg/g (non-DM); 5/50/500 mg/g (DM). Abbreviations: ACR, urine albumin-creatinine ratio; BMI, body mass index; CVD, cardiovascular disease; DM, diabetes mellitus.

*Each column contains 64 dots representing 64 hypothetical scenarios. The dots are shaded from light to dark based on the number of risk factors present, scaled from 0 to 4 based on the presence or absence of CVD, smoking, hypertension, and BMI 35 kg/m².