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1 Deep Brain Stimulation (DBS) and Motor Cortex Stimulation (MCS) for Central Post-Stroke 2 Pain: A Systematic Review And Meta-Analysis Siddarth Kannan¹, Conor S Gillespie², Jeremy Hanemaaijer^{3,4,5}, John Eraifej^{4,5}, Andrew F 3 Alalade^{1,6}, Alex Green^{4,5} 4 ¹ School of Medicine, University of Central Lancashire, Preston, United Kingdom 5 ² Department of Clinical Neurosciences, University of Cambridge, Cambridge, UK 6 7 ³ Department of Neurosurgery, RadboudUMC, Nijmegen, The Netherlands 8 ⁴Oxford Functional Neurosurgery Group, John Radcliffe Hospital, Oxford, United Kingdom 9 ⁵Nuffield Department of Surgical Sciences, University of Oxford, Oxford, United Kingdom 10 ⁶Department of Neurosurgery, Royal Preston Hospital, UK 11 12 Total Number of Pages: 32 13 Total Number of Figures and Tables: 8 14 Corresponding author: Siddarth Kannan 15 School of Medicine 16 17 135A Adelphi St, Preston 18 PR1 7BH Email address: SKannan@uclan.ac.uk 19 20 Telephone: (+44) 7788265067 21 Fax: n/a 22 ORCID: 0009-0004-8703-4859 23 24 25

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1 Abstract

Introduction: Deep Brain Stimulation (DBS) and Motor Cortex stimulation (MCS) are
invasive interventions in order to treat various neuropathic pain syndromes such as Central
Post-Stroke Pain. While each treatment has varying degree of success, comparative analysis
has not yet been performed, and the success rates of these techniques using validated,
objective pain scores have not been synthesised.

Methods: A systematic review and meta-analysis was conducted in accordance with PRISMA
guidelines. Three databases were searched, and articles published from January 2000 October
2024 were included (last search date 25 October 2024). Meta-Analysis was performed using
random effects models. We evaluated the performance of DBS or MCS by assessing studies
that reported pain relief using Visual Analogue Scale (VAS) or Numerical Rating Scale
(NRS) scores.

Results: Of the 478 articles identified, 32 were included in the analysis (330 patients- 139
DBS, & 191 MCS). The improvement in mean VAS score for patients that underwent DBS
post-surgery was 48.6% compared to a score of 53.1% for patients that had MCS. The pooled
number of patients who improved after DBS was 0.62 (95% CI, 0.51-0.71, I2=16%). The
pooled number of patients who improved after MCS was 0.64 (95% CI, 0.53-0.74, I2=40%).

Conclusion: The use of neurosurgical interventions such as DBS and MCS are last-resort
treatments for Central Post-Stroke Pain, with limited studies exploring and comparing these
two techniques. While our study shows that MCS might be a slightly better treatment option,
further research would need to be done to determine the appropriate surgical intervention in
the treatment of Central Post-Stroke Pain.

INTRODUCTION Central Post-Stroke Pain (CPSP) is one of the most challenging and distressing complications that stroke survivors face during recovery, affecting approximately 10-39% of stroke patients [12,18]. The onset of CPSP typically occurs within 1 to 3 months following a stroke, with the majority of cases manifesting symptoms by 6 months [13]. This condition is characterized by chronic pain resulting from damage to the central nervous system, which severely impacts patients' quality of life. The duration of CPSP can be chronic, lasting for months or even years after the initial stroke event. Studies suggest that CPSP affects approximately 8% to 35% of stroke patients, with pain often persisting long after the stroke[6]. The type of pain in CPSP is typically described as sharp, paroxysmal, and often localized to the hemiplegic side, though some patients report a more diffuse pain experience[63]. Research indicates that stroke survivors with chronic pain often exhibit higher levels of depression and anxiety, which can exacerbate pain perception and complicate treatment[18].

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16 Currently, pharmacological treatments, such as anticonvulsants (e.g., gabapentin), selective
17 serotonin reuptake inhibitors (SSRIs), and antidepressants like amitriptyline, are commonly
18 used to manage CPSP. However, these medications are often associated with significant side
19 effects, especially at higher doses, and many patients are unable to tolerate these treatments
20 [27]. As a result, there is a growing need for alternative therapeutic options.

In recent years, invasive neuromodulation techniques have emerged as promising alternatives
to manage CPSP, with Deep Brain Stimulation (DBS) and Motor Cortex Stimulation (MCS)
at the forefront of innovative interventions [17]. Both techniques involve the application of
electrical impulses to specific brain areas, aiming to modulate neural circuitry and provide

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1	relief from chronic pain. However, the targeting areas for both interventions are different. In
2	MCS, two electrode leads are placed over the motor and sensory cortices, while in DBS
3	surgery, deeper located brain structures are targeted, such as the ventral posterolateral (VPL)
4	and ventral posteromedial (VPM) nuclei of the thalamus , the periventricular and
5	periaqueductal grey matter (PVG/PAG), or the rostral anterior cingulate cortex (ACC) [15].
6	While some studies have showcased this potential benefit, pooled comparative analysis has
7	not yet been performed and the success rates of these techniques using validated, objective
8	pain scores have not been synthesised. In this systematic review and meta-analysis, we aim to
9	analyse the effect on pain relief offered by MCS and DBS on patients with Central Post-
10	Stroke Pain using clearly defined outcomes.
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18	MATERIAL AND METHODS
19	Search strategy and selection criteria
20	We conducted this systematic review and meta-analysis according to the Preferred Reporting
21	Items for Systematic Reviews and Meta-Analyses (PRISMA) Guidelines [55].

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We searched PubMed, Embase and Medline database of systematic reviews for full-text articles published in English (Search date 25th October2024). Search terms used a combination of the terms 'Central Post-Stroke Pain', 'Deep Brain Stimulation' and 'Motor Cortex Stimulation', and their associated synonyms. The full search strategy for all databases can be found in Supplementary Tables 1-3. The Population, Intervention, Comparator, Outcome, Study Design (PICOS) criteria was used (Supplementary table 4). Furthermore, excluded reviews and the reference list of retrieved articles were cross-referenced for enriching and completing the included database. We included studies of adults (≥ 18 years) that specifically mentioned the use of either DBS or MCS for the treatment of CPSP. We excluded studies that reported exclusively paediatric populations, and studies that examined other forms of neuropathic pain such as trigeminal neuralgia, diabetic and peripheral neuropathy. We excluded studies that were conference abstracts or if the primary language was not English. Two reviewers (SK, CSG) independently screened titles, abstracts and full texts to include articles. If reviewers failed to reach consensus, a third author was sought for clarification. Data extraction

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Data extraction was completed by two authors independently (SK, CSG). The following data were extracted from included studies: Year published, journal, type of study (Randomized Control Trial [RCT] or observational study), single/multi centre, number of patients with CPSP, and number of these patients that underwent either MCS or DBS, number of patients that saw an improvement in pain, mean postoperative Visual Analogue Scale (VAS) or Numerical Rating Scale (NRS) scores and mean follow up time. The VAS or NRS score was used as primary outcome metric in this study, offering a validated and standardized measure of pain intensity; the scale in each study was measured in 0-10. Any study that did not report individual patient NRS/VAS improvement scores were excluded.

Risk of Bias Assessment

4 Risk of bias assessment was completed by two reviewers independently (SK, CSG).

5 Retrospective studies were classified according to the Newcastle- Ottowa Scale

6 (NOS)[52][69]. NOS is a tool used to assess the quality of non-randomized studies,

7 particularly cohort and case-control studies. Evaluation is based on three broad criteria:

8 Selection, Comparability, and Outcome (for cohort studies) or Exposure (for case-control

9 studies). Each of these criteria includes several sub-criteria, with points awarded to studies

10 based on how well they meet each criterion.

11 Statistical analysis

Baseline characteristics were presented as descriptive frequencies. For meta-analysis, we used random effects models of variables and endpoints, with pooled proportions used for the reduction of pain scores using VAS as a continuous outcome measure. We evaluated the performance of DBS or MCS by assessing studies that reported pain relief using VAS or NRS scores. Pain improvement was defined as a reduction of $\geq 30\%$ on the VAS or NRS score. This is considered clinically significant, as previous studies have indicated that reductions in pain scores of around 30-40% are needed to reflect clinically useful improvements [9,40]. The 30% cut-off for defining improvement was consistently applied across all studies included in this analysis.

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The total number of patients in each study, along with the number of patients who experienced pain improvement of $\geq 30\%$, was extracted. These data were then pooled by calculating proportions representing the percentage of patients experiencing clinically significant pain relief. The pooled proportions were aggregated across studies using random effects models to provide an overall estimate. A 95% confidence interval (CI) was calculated to account for study variability and quantify the uncertainty around the estimated proportions. This approach allowed us to evaluate the clinical effectiveness of the interventions across different studies, providing a measure of precision for the pooled estimates and reflecting the range within which the true proportion is expected to lie 95% of the time. We carried out an additional sensitivity analysis by selectively removing studies at high risk of bias, then re-running the meta-analysis. Data analysis of descriptive statistics was performed using the software Statistical Package for the Social Sciences (version 27; IBM; Armonk; NY). R statistics (Rstudio Version 4.0.1) was used to perform a meta-analysis and create figures, forest, and funnel plots (ggplot2 and meta-packages). For each random effects model, we tested heterogeneity using the maximum restricted likelihood estimator. Prevalence was calculated using pooled proportions methods using the inverse variance method. The I² statistic was used to quantify the percentage of total variation across studies that is due to heterogeneity rather than chance [26]. I² values were interpreted as follows: 0–25%: Low heterogeneity, 26–50%: Moderate heterogeneity, >50%: Significant heterogeneity [25]. Heterogeneity was considered significant when $I^2 > 50\%$ and the P-value < 0.1. In cases of significant heterogeneity, further investigation was conducted to explore the sources of variability among studies. Publication bias was assessed using Egger's test and by inspection of funnel plots.

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3 4 5	1	Sensitivity analysis
5 6 7	2	Further analysis was performed by including studies with a minimum of only 5 patients in
8 9 10	3	order to assess if there was any impact on the overall results.
10 11 12	4	RESULTS
13 14 15 16	5	Study details
10 17 18	6	After removal of duplicates, 97 studies were identified. After full-text assessment, 37 full-text
19 20	7	studies were assessed for inclusion and were finally included, shown in Figure 1
21 22 23	8	(Supplementary table 5). In total, 32 studies were included in the meta-analysis, after
24 25	9	removing five case reports
20 27 28	10	
29 30 31	11	Baseline characteristics
32 33 34	12	The baseline characteristics of included studies are summarized in Table 1 with a total
35 36	13	number of 330 patients. The most common country of published studies was the United
37 38 30	14	Kingdom (26.3%%, n=10). Among the 32 studies included, 16 studies explored the effects of
39 40 41	15	DBS and 16 studies the outcome of MCS. Mean follow up for DBS studies was 18.7 months.
42 43 44	16	For MCS studies, seven studies reported a follow-up time, with a mean of 22.3 months.
45 46 47	17	
48 49 50	18	
50 51 52	19	
53 54 55 56 57 58 59 60	20	Effect of MCS and DBS on VAS pain relief

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1	The analysis of Sixteen MCS studies involving a total of 195 patients
2	[12,14,21,24,27,29,39,44,45,48,61,65,66,71,78,82], revealed that 0.62 (95% CI: [0.53-0.74],
3	$I^2 = 40\%$) experienced pain improvement after undergoing MCS (Figure 2).
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8	Among the sixteen DBS studies included that could be pooled for the meta-analysis
9	[1,7,8,20,22,28,32,37,41,42,46,53,54,57,60,67], the number of patients who have shown an
10	improvement out of each cohort is presented in figure 3. The pooled proportion of patients
11	whose pain scores improved after DBS was 0.62 (95% CI, 0.51-0.71, I ² =16%).
12	
13	Further analysis was conducted based on the stimulation target. Majority of the studies
14	(10/16) targeted the PVG and ventral posterolateral thalamic nucleus (VPL) in 108 patients
15	(Figure 4A). Across these studies, two electrodes are placed, one each in the PVG and VPL
16	unilaterally. The placement of the electrodes is contralateral to the pain affected regions.
17	
18	The remaining studies targeted the anterior cingulate cortex (ACC) in 11 patients , and the
19	posterior limb of the internal capsule (PLIC) in 10 patients. 6 patients in the centromedian-
20	parafascicular nucleus were also included[1]. No significant conclusion regarding the optimal
21	site for pain reduction could be made due to the fewer number of studies and smaller sample
22	size (Figure 4B).

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Sensitivity analysis In DBS, twelve studies met the criteria for sensitivity analysis compared to thirteen MCS studies. Removing four studies (Pereira et al 2007, Kim et al 2012, Hunsche et al 2013 & Son et al 2014) [28,57,67]studies reduced the improvement in pain relief post DBS implantation to 0.59 (95% C1: 0.48-0.69, p=0.24) from 0.624 (95% CI: 0.51-0.71, p=0.27). This was similar for the MCS cohort where the removal of three studies (Yamamoto et al 2007, Delavallée et al 2008 & Ngyuen et al 2008) [12,45,78], reduced the improvement in pain relief to 0.63 (95%) C1: 0.52-0.72, p<0.05) from 0.64 (95% CI: 0.53-0.74, p=0.05) (Supplementary Figure 1-2). Egger's test The potential for publication bias in this meta-analysis was assessed using a funnel plot and Egger's regression test. The funnel plot demonstrated a symmetrical distribution of study effect sizes around the central line, indicating no obvious visual signs of asymmetry (Figure 5). To statistically evaluate this, Egger's test was performed, yielding a non-significant result (t = 0.81, df = 17, p = 0.4297). This p-value, which is well above the conventional threshold of 0.05, suggests no statistically significant asymmetry in the funnel plot. Consequently, these findings provide no strong evidence of publication bias within the studies included in this analysis. Thus, it can be reasonably concluded that the results of this meta-analysis are unlikely to be influenced by publication bias. Four studies were excluded from the Egger's test due to missing standard error.

1	Risk of Bias
2	The Risk of bias for retrospective cohort studies, using the Newcastle-Ottawa Scale. The
3	mean score for all studies was 7.5 (out of a total maximum score of 9), and 5 studies were
4	classified as high risk of bias (Figure 6).
5	
6	Discussion
7	This systematic review and meta-analysis is the first to pool the effect of MCS and DBS in
8	patients with CPSP. By assessing the 32 studies, we found that MCS has an improvement in
9	pain in 64.3% (123/191) of the patients and 62.5% (87/139) in patients receiving DBS.
10	Patients that underwent MCS had a mean VAS improvement of 53.1% compared to 48.6% in
11	patients that underwent DBS.
12	Several theories about CPSP have been proposed. It is believed that CPSP could be caused by
13	an imbalance between the paleospinothalamic (affective-emotional) and neospinothalamic
14	(sensory-discriminative) pathways. Evidence indicates that impaired spinothalamic tract
15	function is related to the pathogenesis of CPSP[51]. It has been hypothesized that MCS
16	decreases thalamic hyperactivity by inducing corticothalamic connections [75,76]. However,
17	the role of N-methyl-D-aspartate (NMDA) receptors and GABAergic interneurons has also
18	been proposed as an interesting aspect of the potential mechanism in MCS [3,13,38].
19	Repetitive Transcranial Magnetic Stimulation (rTMS) of the motor cortex has shown
20	restoration of intracortical inhibition in neuropathic pain patients, where the degree of pain
21	relief correlates with the amount of restoration of inhibition [35]. Additionally, activation of
22	the endogenous opioid descending system and alterations in the limbic system are described

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as potential mechanisms in MCS [19,58].

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As in MCS, the precise pain-relieving effect of DBS remains incompletely elucidated. The main difference between DBS and MCS is that the effect ensured through stimulation in DBS varies depending on the targeted brain area. Pain improvement in VPL/VPM thalamus DBS could be caused by the alteration of the balance of excitatory and inhibitory neurotransmitters within the pain pathways [2,64], which partly aligns with the potential underlying mechanisms of MCS. Stimulating the PVG and PAG will result in releasing endogenous opioid peptides with a decrease in activity of nociceptive signal-transmitting neurons [2,64]. An alteration of the emotional and cognitive aspects of chronic pain is suggested when stimulating brain targets of the limbic system, such as the ACC [2,64]. The lack of insight in the underlying mechanisms of DBS and MCS, as well as the pathophysiology of CPSP, illustrates the complexity of this pain syndrome and the need for a multidimensional approach in pain modulation therapies. While the literature is limited on the direct comparison of DBS vs MCS on CPSP, several studies have scrutinized its efficacy individually. A study by Owen et al, (2007) examined the effect of DBS on 47 patients with CPSP and found a mean improvement in VAS score of 59%[53]. Studies have found varying results with mean VAS pain relief of between 38.1%-68.4%[16,20,32]. However, determinant factors should be taken into consideration, such as heterogeneity in terms of targeted areas: PVG and VPL were the most common sites with 2 studies targeting the ACC and 1 study each targeting the PLIC. Studies looking at the effect of MCS are limited in comparison with DBS. A study by Zhang et al (2018) looked at the effect of MCS on 16 patients with CPSP and found a mean improvement in VAS score of 42.3% [82]. Similar to patients who underwent DBS surgery, MCS has been shown to improve VAS scores by 40 – 63.8%.[47,71].

A study by Nandi et al. (2002), analysed the use of MCS and VPL/PVG DBS on patients
with CPSP. This study concluded that while MCS offers better pain relief, this varies between

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1	patients and is inconsistent in the long-term outcome [43]. Another study by Katayama et al
2	(2002), found that a greater proportion of MCS patients experienced pain relief compared to
3	ventral caudalis (VC) DBS patient [30]. Both studies indicated that DBS is a simpler
4	procedure and generally better tolerated in patients.
5	Similar results were also found in studies comparing MCS vs DBS in other forms of
6	neuropathic pain. Son et al. (2014), directly analysed MCS and DBS in the same eight
7	patients with chronic intractable neuropathic pain. MCS was successful in reducing pain in
8	6/8 compared to 2/8 in DBS. [68]
9	While our results vary slightly with the current literature on mean VAS score, this could be
10	attributed to various factors such as a larger total population of 330 pooled into the analysis,
11	with other studies varying between 6-47 patients. As previously mentioned, the site of DBS
12	insertion is key and could have played an important role in the heterogeneity of the
13	population.
14	
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16	Clinical and research implications
17	Our results have several implications for research and clinical practise. The slightly better
18	success of MCS, solely based on pain scores, could aid clinicians in determining its
19	appropriate use. However, multiple factors should be taken into account before utilizing MCS
20	over DBS as a last-resort treatment for CPSP, including the risk profile, side effects, and
21	patients' medical history.
22	The practical implications, such as treatment cost, could play a crucial role in determining the
23	most suitable intervention, given the similar success-rates. While there are currently no

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studies directly comparing the cost-effectiveness of DBS and MCS for neuropathic pain, a cost analysis study by Zaghi et al. (2009) indicated that MCS incurs significant initial expenses, estimated at \$42,000.00. After 1-year of follow-up, with monthly visits to assess the parameters and configurations, the total cost of treatment is estimated to be around \$45,600.00[80]. In comparison, an analysis by Bishay et al. (2024) of DBS costs across various disorders, neuropathic pain not included, estimated an inflation- and currency-adjusted mean cost of $40,942.85 \pm 17,987.43$ for total DBS surgery. The initial cost of DBS treatment increases to $47,632.22 \pm 23,067.08$ after 1-year of follow-up [5]. Regardless, this is not a direct in-depth cost-effectiveness analysis between MCS and DBS for CPSP. It provides an impression of the estimated total costs for both interventions, including the first year of intensive follow-up. Follow-up care is essential for optimising treatment outcomes in both interventions. Although the nature and intensity of follow-up can vary significantly within patients. The "trial and error" approach to programming different configurations and parameters in patients treated with MCS or DBS is time consuming and patient specific. Therefore, novel research on investigating connectivity-based predictive models could potentially address these challenges. A personalized approach that optimizes configuration and parameters to target specific brain networks may improve the future application of MCS or DBS in the treatment of CPSP. Generally, DBS is considered more invasive than MCS. This procedure involves the implantation of electrodes into deeper brain regions, while MCS requires electrode placement on the surface of the dura mater. Nevertheless, a craniotomy has been carried out over the Rolandic region for appropriate placement in MCS, whereas a burr hole approach is used in DBS surgery.

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Although MCS and DBS are distinct procedures, adverse events appear to be rare and manageable. Potential adverse effects following DBS and MCS surgery include seizures, hematomas, infections, headaches and hardware malfunction. Additionally, some complications are procedure specific, such as epidural fibrosis, electrode migration, and effusion formation are associated with MCS surgery [59]. In DBS, the side effects are more location dependent, such as paraesthesias, muscle spasms, and phosphenes. Side effects could also occur at accustomed therapeutic voltages, the electrode leads therefore should be repositioned into a slightly altered location [70]. The aforementioned findings highlight that while DBS and MCS are effective therapies for chronic neuropathic pain, careful and accurate post-surgical management is not only required for optimizing result, but also cost-effectiveness and minimizing the risk of adverse events. In contrast to the invasive nature of DBS and MCS, non-invasive brain stimulation (NIBS) techniques have been investigated to adjust the excitability of specific functional brain regions[31,72]. The most prevalent utilized NIBS techniques in clinical setting are transcranial magnetic stimulation (TMS) and transcranial direct current stimulation (tDCS). Research indicates that rTMS can effectively reduce neuropathic pain. A study demonstrated that rTMS targeting the motor cortex resulted in significant pain relief for patients suffering from refractory neuropathic pain[74]. The mechanism described by which rTMS alleviates pain is thought to involve modulation of cortical excitability and restoration of normal brain function in pain processing pathways[56]. A clinical trial found that anodal tDCS applied to the motor cortex significantly ameliorated chronic pain and reduced intracortical inhibition, suggesting a potential mechanism for its analgesic effects[4]. While there is no study directly comparing the effect of NIBS on CPSP, a recent meta-analysis on sensory function recovery in stroke patients showed that both tDCS and rTMS significantly outperformed control

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conditions. NIBS was beneficial in the acute and subacute phases of stroke, while a moderate effect was observed in chronic stroke patients [11].

Although rTMS currently has been described as a predictive factor for MCS, the application of NIBS treatments could potentially being integrated in the treatment of CPSP after warrant future research. Moreover, the integration of individualized treatment protocols is likely to play a crucial role in the future of non-invasive neuromodulation[36]. Current research suggests that the same stimulation parameters may not yield uniform effects across different individuals due to variations in brain anatomy and neurophysiology [77]. Personalized approaches, such as adjusting stimulation intensity and targeting specific brain regions based on individual patient profiles, could enhance treatment outcomes and minimize side effects[10].

Though we analyse the efficacy of DBS and MCS on patient outcome there are certain aspects that need to be addressed. The success of MCS could be due to fewer studies present on this topic and the impact of information bias would need to be considered. Various factors impacting pain relief such as severity and location of stroke, age, and social habits could have an impact on the overall outcome [23,50]. Studies have also shown that coping strategies such as social support can influence the effect of pain relief [62,79]. A cross-sectional study found that access to a trusted healthcare professional, living with pain for ≥ 10 years and polypharmacy had a significant effect on the amount of pain relief [81].

Limitations

This study has several limitations. Firstly, all studies included were retrospective, precluding pooled analysis of prospective studies. In addition, the region of the brain where DBS was performed is heterogeneous and could have impacted the effect of pain relief. In MCS, the

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4	1	variety of surgical approaches over time and the exact location of the electrode leads could
5 6 7	2	influence the level of pain relief due to the lack of a standardised protocol. In this meta-
, 8 9	3	analysis, only pain scores have been considered, whereas in pain research, Quality of Life
10 11	4	(QoL) and medication use is at least as important. We also excluded full-text papers not
12 13	5	available in English, restricting paper eligibility.
14 15 16 17	6	
18 19 20	7	Conclusion
21 22	8	The use of neurosurgical interventions, such as DBS and MCS, are a propitious field for the
23 24 25	9	treatment of CPSP, with limited studies exploring and comparing these two techniques.
26 27 28	10	While our study suggests a modest improvement in pain scores with MCS over DBS in
29 30	11	patients suffering from CPSP, these findings should be viewed as preliminary. Pain scores
31 32	12	alone is insufficient to establish MCS as a definitively superior or non-inferior treatment
33 34 35	13	option compared to DBS. Further factors such as cost-effectiveness in pain treatment, long-
36 37	14	term efficacy and multi-dimensional functional outcomes would need to be assessed in order
38 39 40	15	to determine the appropriate surgical neuromodulation for CPSP.
41 42	16	
43 44 45 46	17	Funding: This research received no external funding.
47 48	18	Data Availability Statement: The original contributions presented in the study are included in
49 50 51	19	the article, further inquiries can be directed to the corresponding author.
52 53	20	Conflicts of Interest: The authors declare no conflicts of interest.
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54 55 56	36	Table	legends
57 58 59 60	37	Table	1: Baseline characteristics

1	Figure Legends
2	Figure 1. PRISMA Flow diagram, of study selection for inclusion in this review and meta-
3	analysis.
4	Figure 2: Forest plot showcasing VAS improvement after MCS using and random effects
5	models.
6	Figure 3: Forest plot showcasing VAS improvement after DBS using random effects models.
7	Figure 4 A: Forest plot showcasing VAS improvement in studies targeting the PVG/VPL
8	using random effects models.
9	Figure 4B: Forest plot showcasing VAS improvement in various other targets using random
10	effects models.
11	Figure 5: Funnel plot assessing publication bias.
12	Figure 6: Risk of Bias using the Newcastle-Ottowa Scale
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14	
15	
	1 2 3 4 5 6 7 8 9 10 11 12 13 14 15

10 (26.3%)
5 (13.1%)
4 (10.5%)
3 (7.9%)
3 (7.9%)
2 (5.3%)
2 (5.3%)
2 (5.3%)
1 (2.6%)
1 (2.6%)
1 (2.6%)
1 (2.6%)
1 (2.6%)
1 (2.6%)
33 (100%)
330
210
139
87 (62.5%)

Mean VAS score improvement post-surgery	48.6% (±13.42%)
Mean follow up time in months (Standard	18.67 (±13.52)
deviation)	
MCS	
Patients	191
Number of patients showed improvement	123 (64.3%)
in pain relief	
Mean VAS score improvement post-surgery	53.17% (±9.27%)
Mean follow up time in months (Standard	22.3 (±10.1)
deviation)	



Figure 1: PRISMA flowchart detailing the selection process of studies included in this systematic review and meta-analysis.

53x56mm (300 x 300 DPI)

Study	Events	Total				F	Proportion	95%-CI	Weight (common)	Weight (random)
Zhang 2018	9	16					0.56	[0.30; 0.80]	10.5%	9.5%
Isagulyan 2015	14	20	-				0.70	[0.46; 0.88]	11.2%	9.8%
Fagundes-Pereyra 2010	9	10					0.90	[0.55; 1.00]	2.4%	3.7%
Rasche 2006	3	7					0.43	[0.10; 0.82]	4.6%	5.9%
Tanei 2011	7	8	-		-		0.88	[0.47; 1.00]	2.3%	3.6%
Sokal 2015	12	14			10		0.86	[0.57; 0.98]	4.6%	5.9%
Henssen 2018	7	18					0.39	[0.17; 0.64]	11.4%	9.8%
Sokal 2019	4	6					0.67	[0.22; 0.96]	3.5%	4.9%
Nguyen 2000	10	13	-			_	0.77	[0.46; 0.95]	6.1%	7.1%
Nuti et al 2005	12	22			_		0.55	[0.32; 0.76]	14.5%	10.9%
Hosomi 2008	6	18					0.33	[0.13; 0.59]	10.6%	9.5%
Maarrawi 2013	10	10					1.00	[0.69; 1.00]	1.3%	2.2%
Guo et al 2022	13	21					0.62	[0.38; 0.82]	13.1%	10.5%
Yamamoto 2007	2	2				-	1.00	[0.16; 1.00]	1.1%	1.9%
Delavallee 2008	3	3					1.00	[0.29; 1.00]	1.2%	2.0%
Nguyen 2008	2	3					0.67	[0.09; 0.99]	1.8%	2.9%
Common effect model		191		\Leftrightarrow			0.61	[0.53; 0.68]	100.0%	
Random effects model	- 0 2880	n = 0.0	5	\rightarrow			0.64	[0.53; 0.74]		100.0%
Helei Ugeneily. $T = 40\%$, t	- 0.2009	, <i>μ</i> = 0.0	0.2 0.4	0.6	0.8	1				

Figure 2: Forest plot demonstrating the pooled effect size for improvement in Visual Analog Scale (VAS) scores following Motor Cortex Stimulation (MCS), analysed using random effects models to account for between-study variability.

77x40mm (300 x 300 DPI)

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Study	Events	Total		Proportion	95%-CI	Weight (common)	Weight (random)
Nandi 2003	4	5		0.80	[0.28; 0.99]	2.8%	3.3%
Nandi & Aziz 2004	9	14		0.64	[0.35; 0.87]	11.3%	10.7%
Hamani 2006	5	9		0.56	[0.21; 0.86]	7.8%	8.0%
Rasche 2006	2	11 -		0.18	[0.02; 0.52]	5.7%	6.2%
Owen 2006	12	15		0.80	[0.52; 0.96]	8.4%	8.5%
Pereira 2007	2	2		1.00	[0.16; 1.00]	1.5%	1.8%
Owen 2007	12	18		0.67	[0.41; 0.87]	14.0%	12.5%
Kim 2012	3	3		1.00	[0.29; 1.00]	1.5%	1.9%
Boccard 2013	16	23		0.70	[0.47; 0.87]	17.1%	14.3%
Hunsche 2013	3	3		1.00	[0.29; 1.00]	1.5%	1.9%
Gray 2014	3	5		0.60	[0.15; 0.95]	4.2%	4.8%
Boccard 2014	3	5	_	0.60	[0.15; 0.95]	4.2%	4.8%
Son 2014	4	4		1.00	[0.40; 1.00]	1.6%	1.9%
Levi 2019	2	5 -		0.40	[0.05; 0.85]	4.2%	4.8%
Abdallat 2021	4	11		0.36	[0.11; 0.69]	8.9%	8.9%
Nowacki 2023	3	6		0.50	[0.12; 0.88]	5.3%	5.8%
Common effect model		139		0.62	[0.53; 0.70]	100.0%	
Random effects model			\sim	0.62	[0.51; 0.71]		100.0%
Heterogeneity: $I^2 = 16\%$, τ^2	² = 0.1089	, p = 0.27	0.2 0.4 0.6 0.8	1 1			

Figure 3: Forest plot summarizing the improvement in Visual Analog Scale (VAS) scores achieved after Deep Brain Stimulation (DBS), with random effects modelling to address heterogeneity among studies.

79x42mm (300 x 300 DPI)

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9					Weight Weight	
10	Study	Events Total	Pro	portion 95%-CI (o	common) (random)	
11	Nandi et al 2003 Nandi & Aziz, 2004	4 5 — 9 14	*	0.80 [0.28; 0.99]	3.6% 4.5% 14.7% 14.2%	
12	Hamani et al 2006	5 9 —		0.56 [0.21; 0.86]	10.1% 10.8%	
13	Owen et al 2006	12 15		0.18 [0.02; 0.52]	10.9% 11.4%	
14	Pereira et al 2007 Owen et al 2008	2 2 —— 12 18		1.00 [0.16; 1.00] 0.67 [0.41; 0.87]	1.9% 2.5% 18.2% 16.5%	
16	Boccard et al 2013 Son 2014	16 23 4 4		0.70 [0.47; 0.87] 1.00 [0.40; 1.00]	22.2% 18.8% 2.1% 2.6%	
17	Abdallat 2021 Nowacki 2023	2 2 <u>—</u> 3 6 —		1.00 [0.16; 1.00] 0.50 [0.12; 0.88]	1.9% 2.5% 6.8% 7.8%	
18	Common effect model	109		0.65 [0.55: 0.74]	100.0%	
19	Random effects model	² = 0.4040 = = 0.00		0.65 [0.53; 0.75]	100.0%	
20	Heterogeneity: $I^{-} = 19\%$, τ^{-}	= 0.1248, p = 0.26 0.2	0.4 0.6 0.8 1			
21						
22						
23						
Figure 4A:	Forest plot high	lighting the impr	ovement in VAS	scores observ	ed in studies	focusing on
25 stimulation of 26	f the periventric	ular grey (PVG) a offects models t	and ventroposter	rolateral nuclei	us (VPL), ana t size	lysed through
27	ranuom	enects models t			t size.	
28		79x43m	im (300 x 300 D	PI)		
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Weight Weight

7.3% 7.3%

20.1%

20.1%

20.1%

25.1%

100.0%

95%-CI (common) (random)

7.3% 7.3%

20.1%

20.1%

20.1%

25.1%

100.0%

Proportion

1.00 [0.29; 1.00] 1.00 [0.29; 1.00] 0.60 [0.15; 0.95]

0.60 [0.15; 0.95]

0.40 [0.05; 0.85]

0.50 [0.12; 0.88]

0.59 [0.39; 0.76]

0.59 [0.39; 0.76]



Study

Kim 2012

Gray 2014

Boccard

Levi 2019

Hunsche 2013

Abdallat 2021

Common effect model

Random effects model Heterogeneity: $I^2 = 0\%$, $\tau^2 = 0$, p = 0.68

Events Total

3 3

3 3

3

3 5

2 3 5

5

6

27

0.2 0.4 0.6 0.8 1

Figure 4B: Forest plot presenting the pooled VAS improvements from studies targeting alternative brain

regions, evaluated using random effects models to incorporate variability across studies.

79x43mm (300 x 300 DPI)

50 59







Representatives · Outcome not present at start of study Length of follow-up Domain Confidence Domain No Yes Comparability -Assessment of outcome Ascertainment of data -Adequacy of follow up 25 50 75 ò 100 Percentage (%)

Figure 6: Risk of bias evaluation for included studies, conducted using the Newcastle-Ottawa Scale to assess methodological quality and validity.

79x37mm (300 x 300 DPI)